# Endpoint fluctuations for directed polymers

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We discuss the work [Bol89], primarily following the presentation of [Com17, Chapter 3].

#### 1 Directed polymer model

Let  $\pi = (\pi(t))_{t \in \mathbb{Z}_{\geq 0}}$  be the simple symmetric random walk on  $\mathbb{Z}^d$  started at 0:

$$\pi(t) = \sum_{s=1}^{t} X_s, \qquad X_s \stackrel{\text{i.i.d.}}{\sim} \mathsf{Uniform}(\{\pm e_1, \dots, \pm e_d\})$$

We denote by  $P_{\text{RW}}$  the law of  $\pi$ . This is a probability measure on the set  $\Pi$  of all possible trajectories for  $\pi$ . By the central limit theorem, if  $\pi \sim P_{\text{RW}}$ , then  $\pi(n)$  converges in the diffusive scaling limit to a vector of i.i.d. Gaussians:

$$\lim_{n \to \infty} P_{\text{RW}}\left(\frac{\pi(n)}{\sqrt{n}} \in A\right) = \text{Prob}\big(\mathsf{N}(0, d^{-1}I) \in A\big), \qquad A \subset \mathbb{R}^d$$

Let  $\omega = (\omega(t, x))_{(t,x) \in \mathbb{Z}_{\geq 0} \times \mathbb{Z}^d}$  be a family of i.i.d. Rademacher random variables:  $\mathbb{P}(\omega(t, x) = 1) = \mathbb{P}(\omega(t, x) = -1) = 1/2$ . We denote by  $\mathbb{E}$  the expectation with respect to  $\mathbb{P}$ . The family  $\omega$  is called the **environment**. For  $\beta \in [0, 1]$  and  $n \in \mathbb{Z}_{\geq 0}$ , we define the **Hamiltonian** (or **energy**) of a trajectory  $\pi \in \Pi$  by

$$H_n^{\omega}(\pi) := \prod_{t=1}^n (1 + \beta \omega(t, \pi(t))).$$

The **polymer measure** is the (random) probability measure  $\mu_n^{\omega}$  on  $\Pi$  with Radon–Nikodym derivative

$$\frac{d\mu_n^{\omega}}{dP_{\rm RW}}(\pi) := \frac{H_n^{\omega}(\pi)}{\langle H_n^{\omega} \rangle},$$

where the angle brackets denote expectation with respect to  $P_{\rm RW}$ :

$$\langle H_n^\omega \rangle := \int_{\Pi} H_n^\omega(\pi) \, dP_{\rm RW}(\pi).$$

Example 1.1. If  $\beta = 0$  then  $H_n^{\omega}(\pi) = 1$  for any  $\omega, n, \pi$ , and therefore  $\mu_n^{\omega} = P_{\text{RW}}$ .

It is of great physical importance to understand how and whether the presence of disorder (i.e.  $\beta > 0$ ) affects the geometry and statistics of the simple random walk.

### **2** Gaussian fluctuations in dimension $d \ge 3$

**Theorem 2.1** ([Bol89]). There exists  $\beta_0 > 0$  such that if  $\beta \in (0, \beta_0)$  and  $d \ge 3$ , then

$$\lim_{n \to \infty} \mu_n^{\omega} \left( \frac{\pi(n)}{\sqrt{n}} \in A \right) = P\left( \mathsf{N}(0, d^{-1}I) \in A \right) \quad \text{for } \mathbb{P}\text{-a.e. } \omega.$$

**Corollary 2.2** ([IS88; Bol89]). There exists  $\beta_0 > 0$  such that if  $\beta \in (0, \beta_0)$  and  $d \ge 3$ , then

$$\mathbb{E}_{\mu_n^{\omega}}\left[|\pi(n)|^2\right] = \frac{1}{\langle H_n^{\omega} \rangle} \left\langle |\pi(n)|^2 H_n^{\omega}(\pi) \right\rangle \xrightarrow{n \to \infty} 1 \quad for \ \mathbb{P}\text{-}a.e. \ \omega.$$

We will deduce Theorem 2.1 from the following convergence of moments:

**Theorem 2.3** ([Bol89]). There exists  $\beta_0 > 0$  such that if  $\beta \in (0, \beta_0)$  and  $d \ge 3$ , then for all multiindices  $\alpha \in \mathbb{Z}_{\ge 0}^d$ ,

$$\lim_{n \to \infty} \frac{1}{\langle H_n^{\omega} \rangle} \left\langle \prod_{i=1}^d \left( \frac{\pi_i(n)}{\sqrt{n}} \right)^{\alpha_i} H_n^{\omega}(\pi) \right\rangle = E\left[ \mathsf{N}(0, d^{-1}I)^{\alpha} \right] \quad \text{for } \mathbb{P}\text{-a.e. } \omega,$$

where we adopt the usual multiindex notation  $x^{\alpha} := \prod_{i=1}^{d} x_i^{\alpha_i}$ .

### 3 Proof

#### 3.1 Convergence of the partition function

Let  $\mathcal{F}_n := \sigma \left( \{ \omega(t, x) : t \le n, x \in \mathbb{Z}^d \} \right).$ 

**Lemma 3.1.** The process  $n \mapsto \langle H_n^{\omega} \rangle$  is a nonnegative  $(\mathcal{F}_n)_{n \in \mathbb{Z}_{\geq 0}}$ -martingale with  $\mathbb{E} \langle H_n^{\omega} \rangle = 1$ .

*Proof.* Using that  $\mathbb{E}[1 + \beta \omega(t, x)] = 1$ , we get

$$\mathbb{E}\left[\langle H_{n}^{\omega}\rangle \mid \mathcal{F}_{n-1}\right] = \frac{1}{(2d)^{n}} \sum_{\substack{0=x_{0}, x_{1}, \dots, x_{n}, \\ \mid x_{t} - x_{t-1} \mid = 1}} \mathbb{E}\left[\prod_{t=1}^{n} (1 + \beta\omega(t, x_{t})) \mid \mathcal{F}_{n-1}\right]$$
$$= \frac{1}{(2d)^{n}} \sum_{\substack{0=x_{0}, x_{1}, \dots, x_{n}, \\ \mid x_{t} - x_{t-1} \mid = 1}} \prod_{t=1}^{n-1} (1 + \beta\omega(t, x_{t}))$$
$$= \frac{1}{(2d)^{n-1}} \sum_{\substack{0=x_{0}, x_{1}, \dots, x_{n-1}, \\ \mid x_{t} - x_{t-1} \mid = 1}} \prod_{t=1}^{n-1} (1 + \beta\omega(t, x_{t}))$$
$$= \langle H_{n-1}^{\omega} \rangle.$$

**Lemma 3.2.** The limit  $Z := \lim_{n \to \infty} \langle H_n^{\omega} \rangle$  exists  $\mathbb{P}$ -a.s. and satisfies  $\mathbb{E}Z = 1$  and  $\mathbb{P}(Z > 0) = 1$ .

*Proof.* The a.s. convergence of  $\langle H_n^{\omega} \rangle$  is a direct application of the martingale convergence theorem, which says that non-negative martingales converge a.s. To upgrade this to convergence in  $L^1$ , we analyze the second moment of  $\langle H_n^{\omega} \rangle$ . Let  $\pi, \pi'$  be two independent simple random walks. Then we have

$$\mathbb{E} \langle H_n^{\omega} \rangle^2 = \mathbb{E} \langle H_n^{\omega}(\pi) \rangle \langle H_n^{\omega}(\pi') \rangle$$
  
=  $\mathbb{E} \langle H_n^{\omega}(\pi) H_n^{\omega}(\pi') \rangle$   
=  $\left\langle \mathbb{E} \prod_{t=1}^n (1 + \beta \omega(t, \pi(t)))(1 + \beta \omega(t, \pi'(t))) \right\rangle$   
=  $\left\langle (1 + \beta^2)^{L_n(\pi, \pi')} \right\rangle$ ,

where  $L_n(\pi, \pi')$  is the intersection local time

$$L_n(\pi, \pi') := \#\{1 \le t \le n : \pi(t) = \pi'(t)\}.$$

We have that  $L_n(\pi, \pi') \leq L(\pi, \pi') := \#\{t \geq 1 : \pi(t) = \pi'(t)\}$ . Observe that  $L(\pi, \pi')$  is equal in distribution to the number of times the (non-simple) random walk  $\tilde{\pi} := \pi - \pi'$  returns to 0. Since  $d \geq 3$ , after each time hitting 0 there is a constant probability p > 0 that  $\tilde{\pi}$  never returns to 0 again. <sup>1</sup> This implies an exponential tail  $P_{\text{RW}}[L(\pi, \pi') > k] = (1 - p)^k$ , and in particular for all sufficiently small  $\beta > 0$  we have

$$\left\langle (1+\beta^2)^{L(\pi,\pi')} \right\rangle < \infty.$$

In particular,  $\sup_n \mathbb{E} \langle H_n^{\omega} \rangle^2 < \infty$ . It follows that  $\langle H_n^{\omega} \rangle \to Z$  in  $L^2$ , hence in  $L^1$ . <sup>2</sup> So  $\mathbb{E}Z = 1$ . This implies that  $\mathbb{P}(Z > 0) > 0$ . It can be shown that  $\{Z > 0\}$  is a tail event, so it follows from Kolmogorov's zero-one law that  $\mathbb{P}(Z > 0) = 1$ .

#### **3.2** A martingale construction for convergence of moments

Recall that we want to show, for any fixed multi-index  $\alpha \in \mathbb{Z}_{>0}^d$ , that

$$\lim_{n \to \infty} \frac{1}{\langle H_n^{\omega} \rangle} \left\langle \left( \frac{\pi(n)}{\sqrt{n}} \right)^{\alpha} H_n^{\omega}(\pi) \right\rangle = \frac{1}{(2\pi)^{d/2}} \int_{\mathbb{R}^d} \left( \frac{u}{\sqrt{d}} \right)^{\alpha} e^{-|u|^2/2} \, \mathrm{d}u \quad \mathbb{P}\text{-a.s.}$$

<sup>1</sup>That p is constant follows from the strong Markov property, so in particular we have the identity  $\mathbb{E}L(\pi,\pi') = \sum_{k\geq 0} P_{\mathrm{RW}}(L(\pi,\pi') > k) = \sum_{k\geq 0} (1-p)^k$ . On the other hand, by CLT,  $\tilde{\pi}(n)$  is (approximately) uniformly distributed in a box of volume  $\Theta(n^{d/2})$ . So we get  $\mathbb{E}L(\pi,\pi') = \sum_{n\geq 1} P_{\mathrm{RW}}(\pi(n)=0) = \Theta\left(\sum_{n\geq 1} n^{-d/2}\right)$ , which converges for  $d\geq 3$ . It follows that p>0.

<sup>2</sup>Let M be a martingale with  $\sup_n \mathbb{E}[M_n^2] < \infty$ . One can show that  $\mathbb{E}[(M_n - M_m)^2] = \sum_{k=m}^{n-1} \mathbb{E}[(M_{k+1} - M_k)^2]$  for any  $n \ge m$ . Setting m = 0 and using that  $\sup_n \mathbb{E}[M_n^2] < \infty$  leads to  $\sum_{k=0}^{\infty} \mathbb{E}[(M_{k+1} - M_k)^2] < \infty$ . We conclude that  $M_n$  is Cauchy—hence convergent—in  $L^2$ .

We have taken care of the factor  $1/\langle H_n^{\omega} \rangle$  So it remains to understand the other factor. To this end, we note the following general fact:

**Lemma 3.3.** Let  $\varphi : \mathbb{Z}_{\geq 0} \times \mathbb{Z}^d \to \mathbb{R}$  be a function such that  $\varphi(n, \pi(n))$  is a martingale with respect to the filtration generated by  $\pi$ , i.e.

$$\langle \varphi(n,\pi(n)) \mid \pi(t) : t \le n-1 \rangle = \varphi(n-1,\pi(n-1)).$$

Then  $\langle \varphi(n, \pi(n)) H_n^{\omega}(\pi) \rangle$  is an  $(\mathcal{F}_n)_{n \in \mathbb{Z}_{\geq 0}}$ -martingale.

Proof.

$$\mathbb{E}\left[\left\langle\varphi(n,\pi(n)) H_{n}^{\omega}(\pi)\right\rangle \mid \mathcal{F}_{n-1}\right] = \left\langle\mathbb{E}\left[\varphi(n,\pi(n)) H_{n}^{\omega}(\pi) \mid \mathcal{F}_{n-1}\right]\right\rangle$$
$$= \left\langle\varphi(n,\pi(n)) H_{n-1}^{\omega}(\pi) \mathbb{E}\left[1 + \beta\omega(n,\pi(n))\right]\right\rangle$$
$$= \left\langle\varphi(n,\pi(n)) H_{n-1}^{\omega}(\pi)\right\rangle$$
$$= \left\langle\left\langle\varphi(n,\pi(n)) \mid \pi(t) : t \le n-1\right\rangle H_{n-1}^{\omega}(\pi)\right\rangle$$
$$= \left\langle\varphi(n,\pi(n)) H_{n-1}^{\omega}(\pi)\right\rangle.$$

Unfortunately, plugging in  $(n, \pi(n))$  to the function  $\varphi : (t, x) \mapsto (x/\sqrt{t})^{\alpha}$  does **not** produce a martingale, and we need to add a correction. We derive this correction as follows.

Recall that the  $\alpha$ -th moment can be accessed by differentiating the moment-generating function:

$$\left\langle H_n^{\omega}(\pi) \,\pi(n)^{\alpha} \right\rangle = \partial_{\lambda}^{\alpha} \left\langle H_n^{\omega}(\pi) \, e^{\lambda \cdot \pi(n)} \right\rangle \left|_{\lambda=0} = \left\langle \left. H_n^{\omega}(\pi) \, \partial_{\lambda}^{\alpha} e^{\lambda \cdot x} \right|_{(\lambda,x)=(0,\pi(n))} \right\rangle,$$

where we write  $\partial_{\lambda}^{\alpha} := \prod_{i=1}^{d} \frac{\partial^{\alpha_{i}}}{\partial \lambda_{i}^{\alpha_{i}}}$ . This leads us to define  $\varphi : \mathbb{Z} \times \mathbb{Z}^{d} \to \mathbb{R}$  by

$$\varphi(t,x) := \partial_{\lambda}^{\alpha} e^{\lambda \cdot x - t\varrho(\lambda)} \bigg|_{\lambda=0},$$

where  $\rho$  is the log-moment-generating function of the increments  $X \sim \mathsf{Uniform}(\{\pm e_1, \ldots, \pm e_d\})$ :

$$\varrho(\lambda) := \log \left\langle e^{\lambda \cdot X} \right\rangle = \log \left( \sum_{i=1}^{d} \frac{e^{\lambda_i} + e^{-\lambda_i}}{2d} \right), \qquad \lambda \in \mathbb{R}^d.$$

The role of the factor  $e^{-t\varrho(\lambda)}$  is explained by the following

**Lemma 3.4.**  $\varphi(n, \pi(n))$  is a martingale with respect to the filtration generated by  $\pi$ .

*Proof.* Since  $\langle e^{\lambda \cdot X_n - \varrho(\lambda)} \rangle = 1$ , we get that

$$\left\langle e^{\lambda \cdot \pi(n) - n\varrho(\lambda)} \mid \pi(t) : t \le n - 1 \right\rangle = e^{\lambda \cdot \pi(n-1) - (n-1)\varrho(\lambda)}.$$

Using this, we obtain

$$\begin{split} \left\langle \varphi(n,\pi(n)) \mid \pi(t) : t \leq n-1 \right\rangle &= \partial_{\lambda}^{\alpha} \left\langle e^{\lambda \cdot \pi(n) - n\varrho(\lambda)} \mid \pi(t) : t \leq n-1 \right\rangle \bigg|_{\lambda=0} \\ &= \varphi(n-1,\pi(n-1)). \end{split}$$

Lemma 3.3 now implies the following

**Corollary 3.5.**  $\langle \varphi(n, \pi(n)) H_n^{\omega}(\pi) \rangle$  is an  $(\mathcal{F}_n)_{n \in \mathbb{Z}_{\geq 0}}$ -martingale.

This is promising, but still not enough. So we analyze  $\varphi(t, x)$  more carefully. Notice that we have the decomposition

$$\varphi(t,x) = x^{\alpha} + \varphi_0(t,x) := x^{\alpha} + \sum_{\substack{j \ge 1, \ \gamma \in \mathbb{Z}^d_{\ge 0}, \\ |\gamma| + 2j \le |\alpha|}} A(j,\gamma) t^j x^{\gamma},$$

for some coefficients  $A(j, \gamma)$ . Here, for a multiindex  $\gamma$  we write  $|\gamma| := \gamma_1 + \cdots + \gamma_d$ . Also, notice the moment-generating function of X approximates that of a Gaussian  $G \sim \mathsf{N}(0, d^{-1}I)$ :

$$\left\langle e^{\lambda \cdot X} \right\rangle = 1 + \frac{|\lambda|^2}{2d} + o(|\lambda|^2) \sim e^{|\lambda|^2/2d} = \left\langle e^{\lambda \cdot G} \right\rangle \quad \text{as } \lambda \to 0,$$

where we abuse notation by writing  $\langle \cdot \rangle$  for the expectation in G. It follows that the function

$$\psi(t,x) := \partial_{\lambda}^{\alpha} e^{\lambda \cdot x - t \log \left\langle e^{\lambda \cdot G} \right\rangle} \bigg|_{\lambda=0} = \partial_{\lambda}^{\alpha} e^{\lambda \cdot x - t \frac{|\lambda|^2}{2d}} \bigg|_{\lambda=0}$$

admits a similar decomposition:

$$\psi(t,x) = x^{\alpha} + \psi_0(t,x) := x^{\alpha} + \sum_{\substack{j \ge 1, \ \gamma \in \mathbb{Z}^d_{\ge 0}, \\ |\gamma| + 2j = |\alpha|}} A(j,\gamma) t^j x^{\gamma}.$$

Notice in particular that  $\varphi$  and  $\psi$  have the same coefficients  $A(j, \gamma)$  for  $|\gamma| + 2j = |\alpha|$ . Now we write

$$\left(\frac{\pi(n)}{\sqrt{n}}\right)^{\alpha} = n^{-|\alpha|/2} \varphi(n, \pi(n))$$
$$- \psi_0(1, \pi(n)/\sqrt{n})$$
$$+ n^{-|\alpha|/2} \left[\psi_0(n, \pi(n)) - \varphi_0(n, \pi(n))\right]$$

and take expectations:

$$\frac{1}{\langle H_n^{\omega} \rangle} \left\langle \left(\frac{\pi(n)}{\sqrt{n}}\right)^{\alpha} H_n^{\omega}(\pi) \right\rangle = \frac{n^{-|\alpha|/2}}{\langle H_n^{\omega} \rangle} \left\langle \varphi(n, \pi(n)) H_n^{\omega}(\pi) \right\rangle - \frac{1}{\langle H_n^{\omega} \rangle} \left\langle \psi_0(1, \pi(n)/\sqrt{n}) H_n^{\omega} \right\rangle + \frac{n^{-|\alpha|/2}}{\langle H_n^{\omega} \rangle} \left\langle \left[\psi_0(n, \pi(n)) - \varphi_0(n, \pi(n))\right] H_n^{\omega}(\pi) \right\rangle.$$

We analyze the above terms via induction on  $|\alpha|$ .

When  $\alpha = 0$  then  $\varphi = \psi = 1$ , so the first term is 1 and the second and third terms are 0. So the above expression is equal to the 0-th moment of a Gaussian.

Suppose we have proven the convergence for all multiindices  $|\alpha| \leq k$ , and let  $|\alpha| = k + 1$ . One can argue along the lines of Lemma 3.2 to show that  $n^{-|\alpha|/2} \langle \varphi(n, \pi(n)) H_n^{\omega}(\pi) \rangle \to 0$  a.s. (for details, see [Bol89, Lemma 4]). Since we have shown already that  $\langle H_n^{\omega} \rangle \to Z > 0$ , we conclude that the first term converges to 0 a.s.

Using the preceding argument and the induction hypothesis, we can further conclude that the third term converges to 0. Indeed, the matching of the coefficients  $A(j,\gamma)$  implies that the difference  $\psi_0(t,x) - \varphi_0(t,x)$  has degree at most  $|\alpha| - 3$ .

Finally, we can repeat all of the preceding analysis with the increments  $X_t$  replaced by i.i.d. Gaussians  $G_t$  to show by induction that the second term above converges to  $\langle G^{\alpha} \rangle$ . More precisely, we can use the fact that

$$\langle \psi(1,G) \rangle = \partial_{\lambda}^{\alpha} \left\langle e^{\lambda \cdot G - \log \mathbb{E}e^{\lambda \cdot G}} \right\rangle \Big|_{\lambda=0} = \partial_{\lambda}^{\alpha} \, 1 \Big|_{\lambda=0} = 0, \qquad \forall \alpha \neq 0$$

to deduce the identity  $\langle \psi_0(1,G) \rangle = \langle G^{\alpha} \rangle$ , and then separately apply the induction hypothesis to establish the convergence

$$\frac{1}{\langle H_n^{\omega} \rangle} \left\langle \psi_0(1, \pi(n) / \sqrt{n}) H_n^{\omega}(\pi) \right\rangle \xrightarrow{\text{a.s.}} \left\langle \psi_0(1, G) \right\rangle.$$

## References

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