

ON SELF-AVOIDING POLYGONS AND WALKS: THE SNAKE METHOD VIA PATTERN FLUCTUATION

ALAN HAMMOND

ABSTRACT. For $d \geq 2$ and $n \in \mathbb{N}$, let W_n denote the uniform law on self-avoiding walks of length n beginning at the origin in the nearest-neighbour integer lattice \mathbb{Z}^d , and write Γ for a W_n -distributed walk. We show that the *closing probability* $W_n(|\Gamma_n| = 1)$ that Γ 's endpoint neighbours the origin is at most $n^{-1/2+o(1)}$ in any dimension $d \geq 2$. The method of proof is a reworking of that in [5], which found a closing probability upper bound of $n^{-1/4+o(1)}$. A key element of the proof is made explicit and called the *snake method*. It is applied to prove the $n^{-1/2+o(1)}$ upper bound by means a technique of *Gaussian pattern fluctuation*.

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1. INTRODUCTION

Self-avoiding walk was introduced in the 1940s by Flory and Orr [8, 22] as a model of a long polymer chain in a system of such chains at very low concentration. It is well known among the basic models of discrete statistical mechanics for posing problems that are simple to state but difficult to solve. Two recent surveys are the lecture notes [2] and [16, Section 3].

1.1. The model. We will denote by \mathbb{N} the set of non-negative integers. Let $d \geq 2$. For $u \in \mathbb{R}^d$, let $\|u\|$ denote the Euclidean norm of u . A *walk* of length $n \in \mathbb{N}$ with $n > 0$ is a map $\gamma : \{0, \dots, n\} \rightarrow \mathbb{Z}^d$ such that $\|\gamma(i+1) - \gamma(i)\| = 1$ for each $i \in \{0, \dots, n-1\}$. An injective walk is called *self-avoiding*. A self-avoiding walk γ of length n is said to close (and to be closing) if $\|\gamma(n)\| = 1$. When the *missing edge* connecting $\gamma(n)$ and $\gamma(0)$ is added, a polygon results.

Definition 1.1. Let $\gamma : \{0, \dots, n-1\} \rightarrow \mathbb{Z}^d$ be a closing self-avoiding walk. For $1 \leq i \leq n-1$, let u_i denote the unordered nearest neighbour edge in \mathbb{Z}^d with endpoints $\gamma(i-1)$ and $\gamma(i)$. Let u_n denote γ 's missing edge, with endpoints $\gamma(n-1)$ and $\gamma(0)$. We call the collection of edges $\{u_i : 1 \leq i \leq n\}$ the *polygon* of γ . A *self-avoiding polygon* in \mathbb{Z}^d is defined to be any polygon of a closing self-avoiding walk in \mathbb{Z}^d . The *polygon's length* is its cardinality.

We will usually omit the adjective self-avoiding in referring to walks and polygons. Recursive and algebraic structure has been used to analyse polygons in such domains as strips, as [3] describes.

Note that the polygon of a closing walk has length that exceeds the walk's by one. Polygons have even length and closing walks, odd.

Let SAW_n denote the set of self-avoiding walks γ of length n that start at 0, i.e., with $\gamma(0) = 0$. We denote by W_n the uniform law on SAW_n . The walk under the law W_n will be denoted by Γ . The *closing probability* is $W_n(\Gamma \text{ closes})$.

In [5], an upper bound on the closing probability of $n^{-1/4+o(1)}$ was proved in general dimension. Without significant modifications, the method used cannot prove an upper bound on this quantity that decays more rapidly than $n^{-1/2+o(1)}$. In this article, we rework the method in order to reach the conclusion that this latter decay can indeed be achieved. The next result, which is this conclusion, is the principal result of the present article.

Theorem 1.2. Let $d \geq 2$. For any $\varepsilon > 0$ and $n \in 2\mathbb{N} + 1$ sufficiently high,

$$W_n(\Gamma \text{ closes}) \leq n^{-1/2+\varepsilon}.$$

Our proof of Theorem 1.2 follows rather closely the method of [5]. (The principal difference is remarked on at the end of Section 4.1.) We take the opportunity, however, to present the method in a general guise, with a view to future applications. Indeed, such an application has already been made, as we explain shortly.

The method in question, which is probabilistic in nature, has two elements: the first is a general sufficient condition for proving closing probability upper bounds, and is here explained in a general framework that we call the *snake method*. In order to apply the snake method to reach an upper bound on the closing probability, such as Theorem 1.2, it is necessary to verify the sufficient condition in the method. This step is undertaken in this article, as it was in [5], by a technique of *Gaussian pattern fluctuation*.

1.2. The snake method via polygon joining. In [10], the second application of the snake method appears. The sufficient condition in the method is not verified by pattern fluctuation but instead by a technique of polygon joining. The main result achieved is now stated.

Theorem 1.3. *Let $d = 2$. For any $\varepsilon > 0$, the bound*

$$W_n(\Gamma \text{ closes}) \leq n^{-4/7+\varepsilon}$$

holds on a set of $n \in 2\mathbb{N} + 1$ of limit supremum density at least $1/1250$.

The decay rate in the upper bound achieved here is stronger than that in Theorem 1.2, but the result is restricted to dimension $d = 2$ and to a positive density set of indices $n \in 2\mathbb{N} + 1$.

1.3. An alternative, combinatorial, approach to the closing probability. Let the walk number c_n equal the cardinality of SAW_n . By equation (1.2.10) of [19], the limit $\lim_{n \in \mathbb{N}} c_n^{1/n}$ exists and is positive and finite; it is called the connective constant and denoted by μ , and we have $c_n \geq \mu^n$.

Define the polygon number p_n to be the number of length n polygons up to translation. By (3.2.9) of [19], $\lim_{n \in 2\mathbb{N}} p_n^{1/n} \in (0, \infty)$ exists and equals μ .

The closing probability may be written in terms of the polygon and walk numbers. There are $2n$ closing walks whose polygon is a given polygon of length n , since there are n choices of missing edge and two of orientation. Thus,

$$W_n(\Gamma \text{ closes}) = \frac{2(n+1)p_{n+1}}{c_n}, \tag{1.1}$$

for any $n \in \mathbb{N}$ (but non-trivially only for odd values of n).

Since $c_n \geq \mu^n$, an upper bound on p_n of the form $n^{-1-\chi}\mu^n$ implies a closing probability upper bound of $n^{-\chi+o(1)}$. In [11], such a bound is achieved when $d = 2$, χ is any given value in $(0, 1/2)$, and with $n \in 2\mathbb{N}$ in a set of full density in the even integers. The method used is a polygon joining technique; as such, two quite different routes to Theorem 1.2 when $d = 2$ (and $n \in \mathbb{N}$ odd is typical) are available.

1.4. Conjectural scaling relation and exponent value prediction. We may hypothesize the existence of exponents θ and ξ such that $p_n = n^{-\theta+o(1)}\mu^n$ for $n \in 2\mathbb{N}$ and $c_n = n^{\xi+o(1)}\mu^n$ for $n \in \mathbb{N}$. Writing $W_n(\Gamma \text{ closes}) = n^{-\psi+o(1)}$, the exponent ψ would then exist via (1.1) and equal $\theta + \xi - 1$. (We might call ψ the ‘closing’ exponent.) It is probably fair to say that the existence of the pair (θ, ξ) is uncontroversial but far from being rigorously established (particularly when d equals two and even more so when it equals three; on results in high dimensions, we comment momentarily). The exponent θ is predicted to satisfy a relation with the Flory exponent ν for mean-squared radius of gyration. The latter exponent is specified by the putative formula $\mathbb{E}_{W_n} \|\Gamma(n)\|^2 = n^{2\nu+o(1)}$, where \mathbb{E}_{W_n} denotes the expectation associated with W_n (and where note that $\Gamma(n)$ is the non-origin endpoint of Γ); in essence, $\|\Gamma(n)\|$ is supposed to be typically of order n^ν . The hyperscaling relation that is expected to hold between θ and ν is $\theta = 1 + d\nu$ where the dimension $d \geq 2$ is arbitrary. In $d = 2$, $\nu = 3/4$ and thus $\theta = 5/2$ is expected. That $\nu = 3/4$ was predicted by the Coulomb gas formalism [20, 21] and then by conformal field theory [6, 7]. Hara and Slade [12, 13] used the lace expansion to show that $\nu = 1/2$ when $d \geq 5$ by demonstrating that, for some constant $D \in (0, \infty)$, $\mathbb{E}_{W_n} \|\Gamma_n\|^2 - Dn$ is $O(n^{-1/4+o(1)})$. This value of ν is anticipated in four dimensions as well, since $\mathbb{E}_{W_n} \|\Gamma_n\|^2$ is expected to grow as $n(\log n)^{1/4}$ (For one article in an extensive recent investigation of Bauerschmidt, Brydges and Slade of the continuous-time weakly self-avoiding walk in $d = 4$, see [1].) We mention also that $\xi = 11/32$ is expected when $d = 2$; in light of the $\theta = 5/2$ prediction and (1.1), $W_n(\Gamma \text{ closes}) = n^{-\psi+o(1)}$ with $\psi = 59/32$ is expected. The $11/32$ value was predicted by Nienhuis in [20] and can also be deduced from calculations concerning SLE $_{8/3}$: see [17, Prediction 5].

1.5. Structure of the paper. After some general notation and definitions in Section 2, we present the general apparatus of the snake method in Section 3 and then use the method via Gaussian pattern fluctuation in Section 4 to prove Theorem 1.2.

2. SOME GENERALITIES

2.1. Notation.

2.1.1. Denoting walk vertices and subpaths. For $i, j \in \mathbb{N}$ with $i \leq j$, we write $[i, j]$ for $\{k \in \mathbb{N} : i \leq k \leq j\}$. For a walk $\gamma : [0, n] \rightarrow \mathbb{Z}^d$ and $j \in [0, n]$, we write γ_j in place of $\gamma(j)$. For $0 \leq i \leq j \leq n$, $\gamma_{[i, j]}$ denotes the subpath $\gamma_{[i, j]} : [i, j] \rightarrow \mathbb{Z}^d$ given by restricting γ .

2.1.2. Path reversal. For $n \in \mathbb{N}$ and a length n walk $\gamma : [0, n] \rightarrow \mathbb{Z}^d$, the reversal $\overleftarrow{\gamma} : [0, n] \rightarrow \mathbb{Z}^d$ of γ is given by $\overleftarrow{\gamma}_j = \gamma_{n-j}$ for $j \in [0, n]$.

2.1.3. Walk length notation. We write $|\gamma| = n$ for the length of any $\gamma \in \text{SAW}_n$.

2.1.4. *Maximal lexicographical vertex.* For any finite $V \subset \mathbb{Z}^d$, we write $\text{NE}(V)$ for the lexicographically maximal vertex in V . The notation is extended to any walk $\gamma : [0, n] \rightarrow \mathbb{Z}^d$ by setting $\text{NE}(\gamma)$ equal to $\text{NE}(V)$ with V equal to the image of γ . We will shortly explain the choice of notation NE .

2.1.5. *The two-part decomposition.* In the snake method, we represent any given walk γ in a *two-part* decomposition. This consists of an ordered pair of walks (γ^1, γ^2) that emanate from a certain common vertex and that are disjoint except at that vertex. The two walks are called the *first part* and the *second part*. To define the decomposition, consider any walk γ of length n . We first mention that the common vertex is chosen to be the lexicographically maximal vertex $\text{NE}(\gamma)$ on the walk. Choosing $j \in [0, n]$ so that $\gamma_j = \text{NE}(\gamma)$, the walk γ begins at γ_0 and approaches $\text{NE}(\gamma)$ along the subwalk $\gamma_{[0,j]}$, and then continues to its endpoint γ_n along the subwalk $\gamma_{[j,n]}$. The reversal $\overleftarrow{\gamma}_{[0,j]}$ of the first walk, and the second walk $\gamma_{[j,n]}$, form a pair of walks that emanate from $\text{NE}(\gamma)$. (When j equals zero or n , one of the walks is the length zero walk that only visits $\text{NE}(\gamma)$; in the other cases, each walk has positive length.) The two walks will be the two elements in the two-part decomposition; all that remains is to order them, to decide which is the first part. Associated to each walk is the list of vertices consecutively visited by the walk $(\gamma_j, \dots, \gamma_0)$ and $(\gamma_j, \dots, \gamma_n)$. These lists may be viewed as elements in $\mathbb{Z}^{d(j+1)}$ and $\mathbb{Z}^{d(n+1-j)}$. The first part, γ^1 , is chosen to be the walk in the pair whose list is lexicographically the larger; the second part, γ^2 , is the other walk.

We use square brackets to indicate the two-part decomposition, writing $\gamma = [\gamma^1, \gamma^2]$.

It is useful to visualise in two dimensions the constructions in our arguments, including this one (see Figure 1). When $d = 2$, we adopt the convention that the lexicographical ordering is specified with the second coordinate having priority over the first. In this way, the lexicographically maximal vertex $\text{NE}(\gamma)$ on a walk γ is the vertex that is at least as northerly as any other, and most easterly among those that share its latitude. The notation $\text{NE}(\gamma)$ is chosen in light of this ‘most north then most east’ rule. The notation is also used in higher dimensions; we feel permitted to emphasise $d = 2$ with the choice of notation because the arguments in this article have almost no dependence on dimension $d \geq 2$.

As a small aid to visualization, it is useful to note that if the first γ^1 part of a two-dimensional walk γ for which $\text{NE}(\gamma) = 0$ has length j , then $\gamma^1 : [0, j] \rightarrow \mathbb{Z}^2$ satisfies

- $\gamma_0^1 = 0$ and $\gamma_1^1 = -e_1$;
- $y(\gamma_i^1) \leq 0$ for all $i \in [0, j]$;
- $\gamma_i^1 \notin \mathbb{N} \times \{0\}$ for any $i \in [1, j]$.

2.1.6. *Polygonal invariance.* The following trivial lemma will play an essential role. It is an important indication as to why polygons can be more tractable than walks.

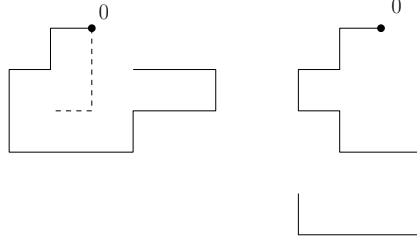


FIGURE 2. *Left:* the bold $\phi \in \text{SAW}_{14}$ and dashed $\gamma \in \text{SAW}_3$ are such that $[\phi, \gamma]$ is a two-part decomposition. Note that $\phi \in \text{First}_{14,14+3} \cap \text{First}_{14,14+4}^c$. *Right:* An element of $\bigcap_{m=1}^{\infty} \text{First}_{14,14+m}$.

We have that

$$\#\{\gamma \in \text{SAW}_n^0 : \gamma \text{ closes}\} = \sum_{j=0}^n \#\{\gamma \in \text{SAW}_n^0 : \gamma \text{ closes}, |\gamma^1| = j\}$$

where, by Lemma 2.1, each term on the right-hand side has equal cardinality. Write $Q = Q_{\delta'} \subseteq \{0, \dots, n\}$ for the index set in the lemma's statement. Note that

$$\begin{aligned} |\text{SAW}_n^0| &\geq |Q| \cdot n^{\alpha' + \delta'} \cdot \frac{1}{n+1} \#\{\gamma \in \text{SAW}_n^0 : \gamma \text{ closes}\} \\ &\geq |Q| \cdot \frac{1}{2} n^{\alpha' - 1 + \delta'} \#\{\gamma \in \text{SAW}_n^0 : \gamma \text{ closes}\}. \end{aligned}$$

Thus, $|Q| \cdot \frac{1}{2} n^{\alpha' - 1 + \delta'} \leq n^{\alpha'}$. \square

2.2.2. Possible first parts and their conditional closing probabilities. Let $d \geq 2$. For $n \in \mathbb{N}$, let $\text{First}_n \subseteq \text{SAW}_n$ denote the set of walks $\gamma : [0, n] \rightarrow \mathbb{Z}^d$ whose lexicographically maximal vertex is $\gamma_0 = 0$. We wish to view First_n as the set of possible first parts of walks $\phi \in \text{SAW}_m^0$ of some length m that is at least n . (In the two-dimensional case, we could be more restrictive in specifying First_n , stipulating if we wish that any element γ satisfies $\gamma_1 = -e_1$. What matters, however, is only that First_n contains all possible first parts.)

Note that, as Figure 2 illustrates, for given $m > n$, only some elements of First_n appear as such first parts, and we now record notation for the set of such elements (whatever the value of $d \geq 2$). Write $\text{First}_{n,m} \subseteq \text{First}_n$ for the set of $\gamma \in \text{First}_n$ for which there exists an element $\phi \in \text{SAW}_{m-n}$ (necessarily with $\text{NE}(\phi) = 0$) such that $[\gamma, \phi]$ is the two-part decomposition of some element $\chi \in \text{SAW}_m$ with $\text{NE}(\chi) = 0$.

In this light, we now define the conditional closing probability

$$q_{n,m} : \text{First}_{n,m} \rightarrow [0, 1], \quad q_{n,m}(\gamma) = W_m^0(\Gamma \text{ closes} \mid \Gamma^1 = \gamma),$$

where here $m, n \in \mathbb{N}$ satisfy $m > n$; note that since $\gamma \in \text{First}_{n,m}$, the event in the conditioning on the right-hand side occurs for some elements of SAW_m , so that the right-hand side is well-defined.

We also identify a set of first parts with *high* conditional closing probability: for $\alpha > 0$, we write

$$\text{HighFirst}_{n,m}^\alpha = \left\{ \gamma \in \text{First}_{n,m} : q_{n,m}(\gamma) > m^{-\alpha} \right\}.$$

2.3. Closing probability lower bound.

Lemma 2.3. *There is a constant $c > 0$ such that, for $n \in 2\mathbb{N} + 1$,*

$$W_n(\Gamma \text{ closes}) \geq \exp \left\{ -cn^{1/2} \right\}.$$

Proof. The Hammersley-Welsh lower bound [9] on the number of self-avoiding *bridges* in terms of walk number is a classical unfolding argument that is recounted in [19, Chapter 3]. It implies that there exists a constant $c_{HW} > 0$ such that, for all $n \in \mathbb{N}$,

$$c_n \leq e^{c_{HW}n^{1/2}} \mu^n, \tag{2.1}$$

where recall that $\mu = \lim_n c_n^{1/n}$. We may now use the bound $p_n \geq e^{-cn^{1/2}} \mu^n$ in [15, Theorem 3] and (1.1) to conclude. Alternatively (but similarly), the lemma is proved from (2.1) by further unfolding in [4]: see Lemma 5 and the proof of Proposition 3 in that paper. \square

3. THE SNAKE METHOD: GENERAL ELEMENTS

In this section, we present in a general form the snake method. The method is a proof-by-contradiction technique which will be applied, alongside other ideas, to prove Theorem 1.2 in the final Section 4. The overall strategy of proof in both sections follows that in [5].

The snake method is used to prove upper bounds on the closing probability, and assumes to the contrary that to some degree this probability has slow decay. For the technique to be used, two ingredients are needed.

- (1) A charming snake is a walk or polygon γ many of whose subpaths beginning at $\text{NE}(\gamma)$ have high conditional closing probability, when extended by some common length. It must be shown that charming snakes are not too atypical.
- (2) A general procedure is then specified in which a charming snake is used to manufacture huge numbers of alternative self-avoiding walks. These alternatives overwhelm the polygons in number and show that the closing probability is very small, contradicting the assumption.

The first step certainly makes use of the assumption of slow decay on the closing probability. It is not however a simple consequence of this assumption. It is this first step that will be carried out in Section 4, by means of a technique of Gaussian pattern fluctuation. (In [10], this step is completed via polygon joining to prove Theorem 1.3.)

In contrast to these two different approaches that are used to implement the first step, the second step is performed using a general tool, valid in any dimension $d \geq 2$, that we present in this section. This step draws inspiration from the notion that reflected walks offer alternatives to closing (or near closing) ones that appears in Madras' derivation [18] of lower bounds on moments of the endpoint distance under W_n .

3.1. The general apparatus of the method.

3.1.1. *Parameters.* The snake method has three exponent parameters:

- the inverse charm $\alpha > 0$;
- the snake length $\beta \in (0, 1]$;
- and the charm deficit $\eta \in (0, \beta)$.

It has two index parameters:

- $n \in 2\mathbb{N} + 1$ and $\ell \in \mathbb{N}$, with $\ell \leq n$.

3.1.2. *Charming snakes.* Here we define these creatures.

Definition 3.1. *Let $\alpha > 0$, $n \in 2\mathbb{N} + 1$, $\ell \in [0, n]$, $\gamma \in \text{First}_{\ell, n}$, and $k \in [0, \ell]$ with $\ell - k \in 2\mathbb{N}$. We say that γ is (α, n, ℓ) -charming at (or for) index k if*

$$W_{k+n-\ell}^0 \left(\Gamma \text{ closes} \mid |\Gamma^1| = k, \Gamma^1 = \gamma_{[0, k]} \right) > n^{-\alpha}. \quad (3.1)$$

(The event that $|\Gamma^1| = k$ in the conditioning is redundant and is recorded for emphasis.) Note that an element $\gamma \in \text{First}_{\ell, n}$ is (α, n, ℓ) -charming at index k if in selecting a length $n - \ell$ walk beginning at 0 uniformly among those choices that form the second part of a walk whose first part is $\gamma_{[0, k]}$, the resulting $(k + n - \ell)$ -length walk closes with probability exceeding $n^{-\alpha}$. (Since we insist that n is odd and that ℓ and k share their parity, the length $k + n - \ell$ is odd; the condition displayed above could not possibly be satisfied if this were not the case.) Note that, for any $\ell \in [0, n]$, $\gamma \in \text{First}_{\ell, n}$ is (α, n, ℓ) -charming at the special choice of index $k = \ell$ precisely when $\gamma \in \text{HighFirst}_{\ell, n}^\alpha$. When $k < \ell$ with $k + n - \ell$ of order n , the condition that $\gamma \in \text{First}_{\ell, n}$ is (α, n, ℓ) -charming at index k is *almost* the same as $\gamma_{[0, k]} \in \text{HighFirst}_{k, k+n-\ell}^\alpha$; (the latter condition would involve replacing n by $k + n - \ell$ in the right-hand side of (3.1)).

For $n \in 2\mathbb{N} + 1$, $\ell \in [0, n]$, $\alpha, \beta > 0$ and $\eta \in (0, \beta)$, define the charming snake set

$$\text{CS}_{\beta, \eta}^{\alpha, \ell, n} = \left\{ \gamma \in \text{First}_{\ell, n} : \gamma \text{ is } (\alpha, n, \ell)\text{-charming} \right. \\ \left. \text{for at least } n^{\beta-\eta}/4 \text{ indices belonging to the interval } [\ell - n^\beta, \ell] \right\}.$$

For any element of $\gamma \in \text{First}_{\ell, n}$, think of an extending snake consisting of $n^\beta + 1$ terms $(\gamma_{[0, \ell - n^\beta]}, \gamma_{[0, \ell - n^\beta + 1]}, \dots, \gamma_{[0, \ell]})$. If $\gamma \in \text{CS}_{\beta, \varepsilon}^{\alpha, \ell, n}$, then there are many charming terms in this snake, for each of which there is a high conditional closing probability for extensions of a *shared* continuation length $n - \ell$.

3.1.3. A general tool for the method's second step. For the snake method to produce results, we must work with a choice of parameters for which $\beta - \eta - \alpha > 0$. (The method could be reworked to handle equality in some cases.) Here we present the general tool for carrying out the method's second step. This technique was already presented in [5, Lemma 5.8], and our treatment differs only by using notation adapted for the snake method with general parameters.

The tool asserts that, if $\beta - \eta - \alpha > 0$ and even a very modest proportion of snakes are charming, then the closing probability drops precipitously.

Theorem 3.2. *Let $d \geq 2$. Set $c = 2^{\frac{1}{5(4d+1)}} > 1$ and set $K = 20(4d + 1)^{\frac{\log(4d)}{\log 2}}$. Suppose that the exponent parameters satisfy $\delta = \beta - \eta - \alpha > 0$. If the index parameter pair (n, ℓ) satisfies $n \geq K^{1/\delta}$ and*

$$\mathbb{P}_{n+1}(\Gamma_{[0, \ell]} \in \text{CS}_{\beta, \eta}^{\alpha, \ell, n}) \geq c^{-n^\delta/2}, \quad (3.2)$$

then

$$W_n(\Gamma \text{ closes}) \leq 2(n+1)c^{-n^\delta/2}.$$

Note that since the closing probability is predicted to have polynomial decay, the hypothesis (3.2) is never satisfied in practice. For this reason, the snake method will always involve argument by contradiction, with (3.2) being established under a hypothesis that the closing probability is, to some degree, decaying slowly.

3.2. A charming snake creates huge numbers of reflected walks. Here is the principal component of the proof of Theorem 3.2.

Proposition 3.3. *Let $d \geq 2$. Set $\delta = \beta - \eta - \alpha$ and suppose that $\delta > 0$. Let $\phi \in \text{CS}_{\beta, \eta}^{\alpha, \ell, n}$. With $c > 1$ and $K > 0$ specified in Theorem 3.2, we have that, if $n \geq K^{1/\delta}$, then*

$$\#\left\{\gamma \in \text{SAW}_n^* : \gamma_{[0, \ell]} = \overleftarrow{\phi}\right\} \geq c^{n^\delta} \cdot \#\left\{\gamma \in \text{SAW}_n^* : \text{NE}(\gamma) = 0, \gamma^1 = \phi\right\}.$$

Note here that walks beginning with the reversal of an element $\phi \in \text{SAW}_\ell$ will necessarily not begin at the origin, and thus we employ the notation introduced in Subsection 2.1.7.

Proof of Proposition 3.3. Let W denote the set of walks γ of length $n - \ell$ that originate at 0 and for which $\text{NE}(\gamma) = 0$. This set is exactly the same as $\text{First}_{n-\ell}$ for all $d \geq 2$. We will be using that W contains all possible length $n - \ell$ walks that

form the second (rather than the first) part of the two-part decomposition of some walk of at least this length, and thus change notation.

Let P denote the uniform measure on the set W . We will denote by Γ a random variable distributed according to P . In particular, Γ is contained in the lower half-space including the origin. (When $d = 2$, we mean the region on or below the x -axis, and, when $d \geq 3$, the region of non-positive e_1 -coordinate.)

We now extend the notion of closing walk by saying that γ' *closes* γ if $\gamma_0 = \gamma'_0$ and the endpoints of γ and γ' are adjacent. We say that γ' *avoids* γ if no vertex except γ_0 is visited by both γ' and γ .

We are given $\phi \in \text{First}_{\ell, n}$ such that $\phi \in \text{CS}_{\beta, \eta}^{\alpha, \ell, n}$. By definition, we may find indices $j_1 < j_2 < \dots < j_{\lceil n^{\beta-\eta}/4 \rceil}$ lying in $[\ell - n^\beta, \ell]$ at each of which ϕ is (α, n, ℓ) -charming.

For $1 \leq i \leq \lceil n^{\beta-\eta}/4 \rceil$, define the events

$$A_i = \left\{ \Gamma \text{ avoids } \phi_{[0, j_i]} \right\} \quad \text{and} \quad C_i = \left\{ \Gamma \text{ closes } \phi_{[0, j_i]} \right\}.$$

Also, define the set $A = \{ \gamma \in W : \gamma \text{ avoids } \phi_{[0, \ell]} \}$.

Since ϕ is (α, n, ℓ) -charming at index j_i ,

$$P\left(\Gamma \text{ closes } \phi_{[0, j_i]} \mid \Gamma \text{ avoids } \phi_{[0, j_i]}\right) = P(C_i \mid A_i) > n^{-\alpha}. \quad (3.3)$$

Write $k = \lceil 4d n^\alpha \rceil$. (Note that $k \leq n^{\beta-\eta}/4$ holds for n high enough since δ is supposed positive.) Any realization $\Gamma \in W$ is in at most $2d$ events C_i . Hence, by (3.3) and the A_j being decreasing,

$$2d \geq \sum_{i=1}^k P(C_i) \geq \sum_{i=1}^k P(C_i \mid A_i) \cdot P(A_k) \geq 4d P(A_k).$$

Therefore, $P(A_k) \leq \frac{1}{2}$. If the procedure is repeated for indices between $k+1$ and $2k$, one obtains

$$2d \geq \sum_{i=k+1}^{2k} P(C_i \mid A_k) \geq \sum_{i=k+1}^{2k} P(C_i \mid A_i) \cdot P(A_{2k} \mid A_k) \geq 4d P(A_{2k} \mid A_k),$$

and thus $P(A_{2k} \mid A_k) \leq 1/2$. Since $A_{2k} \subset A_k$, we find

$$P(A_{2k}) = P(A_k)P(A_{2k} \mid A_k) \leq \frac{1}{4}.$$

In these inequalities, we see the powerful bootstrap mechanism at the heart of the snake method, demonstrating that $P(A_{(i+1)k})$ is at most one-half of $P(A_{ik})$. The mechanism works because the method's definitions imply that all walk extensions are of common length $n - \ell$, and the avoidance conditions are monotone (i.e., the events A_i are decreasing).

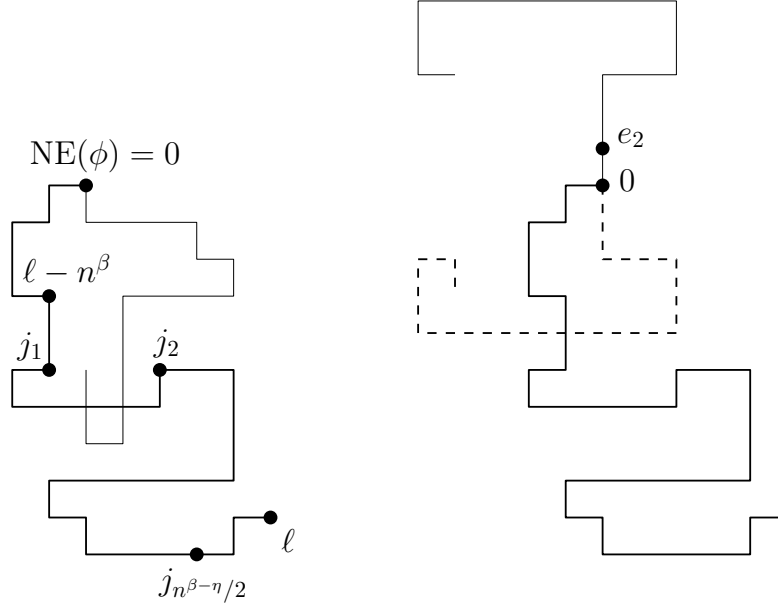


FIGURE 3. On the left, $\phi \in \text{First}_{\ell, n}$ is depicted in bold. The labels along ϕ are indices. A walk in $W \cap A_1 \cap C_1 \cap A_2^c$ is also shown. On the right, another element $\gamma \in W$ is depicted with dashed lines. The output of the three-part concatenation procedure used at the end of the proof of Proposition 3.3 is shown above 0.

Indeed, the procedure may be repeated $\lfloor \frac{n^{\beta-\eta}}{4k} \rfloor \geq \frac{n^{\beta-\eta-\alpha}}{4(4d+1)} - 1$ times. Recalling that $\phi = \phi_{[0, \ell]}$, we obtain

$$\frac{|A|}{|W|} = P(\Gamma \text{ avoids } \phi) \leq P(A_{\lceil n^{\beta-\eta}/4 \rceil}) \leq 2^{-\lfloor \frac{n^{\beta-\eta}}{4k} \rfloor} \leq 2^{1 - \frac{n^{\beta-\eta-\alpha}}{4(4d+1)}},$$

whatever the value of $n \in 2\mathbb{N} + 1$.

The set A contains all length $n - \ell$ walks γ for which $[\phi, \gamma]$ is the two-part decomposition of a walk of length n with $\text{NE} = 0$. Thus,

$$|A| = \#\left\{ \gamma \in \text{SAW}_n^* : \text{NE}(\gamma) = 0, |\gamma^1| = \ell, \gamma^1 = \phi \right\}.$$

On the other hand, for $\gamma \in W$, consider the walk obtained by concatenating three paths (and illustrated in Figure 3). When $d = 2$, these are: the reversal $\overleftarrow{\phi}$ of ϕ ; the edge e_2 ; and the e_2 -translation of the reflection of γ in the horizontal axis. When $d \geq 3$, we substitute e_1 for e_2 , and the zero-coordinate hyperplane in the e_1 -direction for the horizontal axis, to specify these paths. The walk that results has length $n + 1$ and is self-avoiding. By deleting the last edge of such walks, we obtain at least $|W|/2d$ walks of length n , each of which follows $\overleftarrow{\phi}$ in its first ℓ

steps. Thus,

$$\#\{\gamma \in \text{SAW}_n^* : \gamma_{[0,\ell]} = \overleftarrow{\phi}\} \geq |W|/2d.$$

The three preceding displayed equations combine to prove that the ratio of the cardinalities on the left and right-hand sides in Proposition 3.3 is at least $(2d)^{-1} 2^{\frac{n^{\beta-\eta-\alpha}}{4(4d+1)}-1}$. Noting the lower bound on n stated in Proposition 3.3 completes the proof of this result. \square

Proof of Theorem 3.2. Write $\text{CS} = \text{CS}_{\beta,\eta}^{\alpha,\ell,n}$. Now using Proposition 3.3 in the second and (3.2) in the fourth inequalities,

$$\begin{aligned} |\text{SAW}_n| &\geq \sum_{\phi \in \text{CS}} \#\{\gamma \in \text{SAW}_n : \gamma_{[0,\ell]} = -\phi_\ell + \overleftarrow{\phi}\} \\ &= \sum_{\phi \in \text{CS}} \#\{\gamma \in \text{SAW}_n^* : \gamma_{[0,\ell]} = \overleftarrow{\phi}\} \\ &\geq c^{n^\delta} \cdot \#\{\gamma \in \text{SAW}_n^* : \text{NE}(\gamma) = 0, \gamma^1 \in \text{CS}\} \\ &\geq c^{n^\delta} \cdot \#\{\gamma \in \text{SAW}_n^* : \text{NE}(\gamma) = 0, \gamma^1 \in \text{CS}, \gamma \text{ closes}\} \\ &= c^{n^\delta} \cdot \#\{\gamma \in \text{SAP}_{n+1} : \gamma_{[0,\ell]} \in \text{CS}\} \geq c^{n^\delta/2} \cdot |\text{SAP}_{n+1}|. \end{aligned}$$

Thus, $p_{n+1}/c_n \leq c^{-n^\delta/2}$. We find that

$$W_n(\Gamma \text{ closes}) = 2(n+1)p_{n+1}/c_n \leq 2(n+1)c^{-n^\delta/2}.$$

\square

4. THE SNAKE METHOD APPLIED VIA GAUSSIAN PATTERN FLUCTUATION

Here we prove Theorem 1.2 by assuming that its conclusion fails and seeking a contradiction. By a relabelling of $\varepsilon > 0$, we may express the premise that the conclusion fails in the form that, for some $\varepsilon > 0$ and infinitely many $n \in 2\mathbb{N} + 1$,

$$W_n(\Gamma \text{ closes}) \geq n^{-1/2+4\varepsilon}. \quad (4.1)$$

We fix the three snake method exponent parameters. Fixing a given choice of $\varepsilon \in (0, 1/4)$, we set them equal to $\alpha = 1/2 - 2\varepsilon$, $\beta = 1/2$ and $\eta = 0$. We will argue that the hypothesis (3.2) is comfortably satisfied, with charming snakes being the norm.

Proposition 4.1. *There exists a positive constant C such that, if $n \in 2\mathbb{N} + 1$ and $\varepsilon \in (0, 1/4)$ satisfy (4.1) as well as $n \geq \max\{4^{1/\varepsilon}, C\}$, then there exists $\ell \in [n/4, 3n/4]$ for which*

$$P_{n+1}(\Gamma_{[0,\ell]} \in \text{CS}_{1/2,0}^{1/2-2\varepsilon,\ell,n}) \geq 1 - n^{-\varepsilon/7}.$$

Proof of Theorem 1.2. Given our choice of the three exponent parameters, note that $\delta = \beta - \eta - \alpha$ equals 2ε and is indeed positive. The conclusion of Theorem 3.2 contradicts (4.1) if n is high enough. Thus (4.1) is false for all n sufficiently high. Since $\varepsilon \in (0, 1/4)$ may be chosen arbitrarily small, we are done. \square

It remains only to prove Proposition 4.1, and the rest of the section is devoted to this proof. The plan in outline has two steps. In the first, implemented in Lemma 4.2, we will infer from the closing probability lower bound (4.1) a deduction that for typical indices ℓ close to $n/2$, the initial (or first part) segment $\Gamma_{[0,\ell]}$ of a \mathbb{P}_n -distributed polygon typically has a not-unusually-high conditional closing probability for a second part extension of length $n - \ell$. This step is a straightforward, Fubini-style statement. In the second step, we work with this deduction to build a charming snake. For this, we want an inference of the same type, but with the second part length remaining fixed, even as the first part length varies over an interval of length close to $n^{1/2}$. Beginning in Subsection 4.2, the mechanism of Gaussian fluctuation in local patterns along the polygon is used for this second step. Crudely put, typical deviations in the number of local configurations (type I and type II patterns) in the initial and final ten-percent-length segments of the polygon create an ambiguity in the location marked by the index $\ell \approx n/2$ of order $n^{1/2}$. This square-root fuzziness yields charming snakes.

4.1. Setting the snake method index parameters. We now specify the values of n and ℓ . The value of $n \in 2\mathbb{N} + 1$ is supposed to satisfy (4.1) for our given $\varepsilon \in (0, 1/4)$, as well as the bound $n \geq 4^{1/\varepsilon}$. Applying Lemma 2.2 with $\alpha' = 1/2 - 4\varepsilon$ and $\delta' = \varepsilon$, and noting that this lower bound on n ensures that $2n^{1-\varepsilon} < \#[n/4, 3n/4]$, we find that we may select ℓ to lie in $[n/4, 3n/4]$ and to satisfy

$$\#\{\gamma \in \text{SAW}_n^0 : |\gamma^1| = \ell\} < n^{1/2-3\varepsilon} \cdot \#\{\gamma \in \text{SAW}_n^0 : |\gamma^1| = \ell, \gamma \text{ closes}\},$$

or equivalently

$$\mathbb{W}_n^0(\Gamma \text{ closes} \mid |\Gamma^1| = \ell) > n^{-1/2+3\varepsilon}.$$

The value of ℓ is so fixed henceforth in the proof of Proposition 4.1.

Lemma 4.2.

$$\mathbb{P}_{n+1}(\Gamma_{[0,\ell]} \notin \text{HighFirst}_{\ell,n}^{1/2-2\varepsilon}) \leq n^{-\varepsilon}.$$

Proof. Note that $\Gamma_{[0,\ell]}$ under \mathbb{P}_{n+1} shares its law with the first part Γ^1 under $\mathbb{W}_n^0(\cdot \mid \Gamma \text{ closes}, |\Gamma^1| = \ell)$. For this reason, the statement may be reformulated

$$\mathbb{W}_n^0(\Gamma^1 \notin \text{HighFirst}_{\ell,n}^{1/2-2\varepsilon} \mid |\Gamma^1| = \ell, \Gamma \text{ closes}) \leq n^{-\varepsilon}. \quad (4.2)$$

To derive (4.2), set p equal to its left-hand side. Note that

$$\begin{aligned} & \#\{\gamma \in \text{SAW}_n^0 : |\gamma^1| = \ell\} \\ & \geq \#\{\gamma \in \text{SAW}_n^0 : |\gamma^1| = \ell, \gamma^1 \notin \text{HighFirst}_{\ell,n}^{1/2-2\varepsilon}\} \end{aligned}$$

$$\begin{aligned}
 &= \sum_{\phi \in \text{First}_{\ell,n} \setminus \text{HighFirst}_{\ell,n}^{1/2-2\varepsilon}} \#\{\gamma \in \text{SAW}_n^0 : \gamma^1 = \phi\} \\
 &\geq \sum_{\phi \in \text{First}_{\ell,n} \setminus \text{HighFirst}_{\ell,n}^{1/2-2\varepsilon}} n^{1/2-2\varepsilon} \cdot \#\{\gamma \in \text{SAW}_n^0 : \gamma^1 = \phi, \gamma \text{ closes}\} \\
 &= n^{1/2-2\varepsilon} \cdot \#\{\gamma \in \text{SAW}_n^0 : |\gamma^1| = \ell, \gamma^1 \notin \text{HighFirst}_{\ell,n}^{1/2-2\varepsilon}, \gamma \text{ closes}\} \\
 &= n^{1/2-2\varepsilon} \cdot p \cdot \#\{\gamma \in \text{SAW}_n^0 : |\gamma^1| = \ell, \gamma \text{ closes}\},
 \end{aligned}$$

whence

$$n^{1/2-2\varepsilon} \cdot p \leq W_n^0 \left(\Gamma \text{ closes} \mid |\Gamma^1| = \ell \right)^{-1},$$

whose right-hand side we know to be at most $n^{1/2-3\varepsilon}$. Thus, $p \leq n^{-\varepsilon}$, and we have verified (4.2). \square

Remark. Lemma 4.2 may be compared to [5, Lemma 5.5] with $k = 0$. The former result states that $\text{HighFirst}_{\ell,n}^{1/2-2\varepsilon}$ membership by $\Gamma_{[0,\ell]}$ is the norm under \mathbb{P}_{n+1} , while the latter merely asserts that a comparable membership is not unlikely (having probability at least $n^{-1/4+2\varepsilon}$). It may be possible to improve the inequality in [5, (5.7)] to reflect the fact that the conditioning on closing leading from the law W_n to \mathbb{P}_{n+1} reweights the measure on first parts proportionally in accordance with the conditional closing probability given the first part. Analysing this reweighting may lead to a replacement of the right-hand side of [5, (5.5)] by a term of the form $1 - n^{-o(1)}$ and, alongside other suitable changes, permit a derivation of Theorem 1.2.

4.2. Patterns and shells. Patterns are local configurations in self-avoiding walks that are the subject of a famous theorem [15] due to Kesten that we will shortly state. For our present purpose, we identify two particular patterns.

Definition 4.3 (Type I/II patterns). *A pair of type I and II patterns is a pair of self-avoiding walks χ^I, χ^{II} , both contained in the cube $[0, 3]^d$, with the properties that*

- χ^I and χ^{II} both visit all vertices of the boundary of $[0, 3]^d$,
- χ^I and χ^{II} both start at $(1, 3, 1, \dots, 1)$ and end at $(2, 3, 1, \dots, 1)$,
- the length of χ^{II} exceeds that of χ^I by two.

Figure 4 depicts examples of such patterns for $d = 2$. The existence of such pairs of walks for any dimension $d \geq 2$ may be easily checked, and no details are given here. Fix a pair of type I and II patterns henceforth.

A pattern χ is said to occur at step k of a walk γ if $\gamma_{[k, k+|\chi|]}$ is a translate of χ (where recall that $|\chi|$ is the length of χ). A *slot* of γ is any translate of $[0, 3]^d$ containing $\gamma_{[k, k+|\chi|]}$ where a pattern χ of type I or II occurs at step k of γ . Note that the slots of γ are pairwise disjoint.

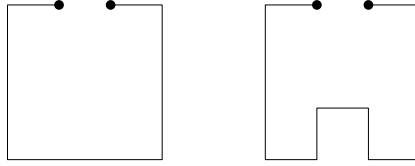


FIGURE 4. An example of type I and II patterns for $d = 2$.

In [5], the notion of shell was introduced. A shell is an equivalence class of self-avoiding walks under the relation that two walks are identified if one may be obtained from the other by changing some patterns of type *I* to be of type *II* and *vice versa*. The walks in a given shell share a common set of slots, but they are of varying lengths. The shell of a given walk γ is denoted $\zeta(\gamma)$.

Consider two walks γ^1 and γ^2 with $\gamma_0^1 = \gamma_0^2$. Recall that γ^2 *avoids* γ^1 if no other vertex is visited by both walks. The next fact is crucial to our reasons for considering shells; its almost trivial proof is omitted.

Lemma 4.4. *For some $m \in \mathbb{N}$, let $\gamma \in \text{SAW}_m$ and let $\gamma' \in \zeta(\gamma)$. A walk beginning at γ_0 avoids γ if and only if it avoids γ' .*

The reader may now wish to view Figure 5 and its caption for an expository overview of the snake method via Gaussian pattern fluctuation. We mention also that this Gaussian fluctuation has been utilized in [14] to prove a $n^{1/2-o(1)}$ lower bound on the absolute value of the writhe of a typical length n polygon.

We will make some use of the notion of shell, but will predominantly consider a slightly different definition, the $(n + 1)$ -local shell, which we now develop. This new notion concerns polygons rather than walks.

Recall that the parameter $n \in 2\mathbb{N} + 1$ has been fixed in Subsection 4.1. (The upcoming definitions do not require that $n \in 2\mathbb{N} + 1$ be fixed in this particular way in order to make sense, but, when we make use of the definitions, it will be for this choice of n .) Define an equivalence relation \sim on SAP_{n+1} as follows. For any $\gamma \in \text{SAP}_{n+1}$, let γ^{empty} denote the polygon in the shell of γ that has no type *II* patterns, (formed by switching every type *II* pattern of γ into a type *I* pattern). Thus, $\gamma^{\text{empty}} \in \text{SAP}_{n+1-2T_{II}(\gamma)}$, where $T_{II}(\gamma)$ denotes the total number of type *II* patterns in γ . A type *II* pattern contains thirteen edges (in $d = 2$; at least this number in higher dimensions), and these patterns are disjoint, so $T_{II}(\gamma) \leq (n + 1)/13$. Thus, the length of γ^{empty} is at least $11(n + 1)/13$. Let S_1 denote the set of slots in γ that are slots in $\gamma_{[0, (n+1)/10]}^{\text{empty}}$, and, writing l^{empty} for the length of γ^{empty} , let S_2 denote the set of slots in γ that are slots in $\gamma_{[l^{\text{empty}} - (n+1)/10, l^{\text{empty}}]}^{\text{empty}}$. Note that S_1 and S_2 are disjoint. We further write $N_I(\gamma)$ and $N_{II}(\gamma)$ for the number of patterns of the given type in the slots $S_1 \cup S_2$, and $N_I^1(\gamma)$ and $N_I^2(\gamma)$ for the number of type *I* patterns occupying slots in S_1 and in S_2 ; and similarly for $N_{II}^1(\gamma)$ and $N_{II}^2(\gamma)$.

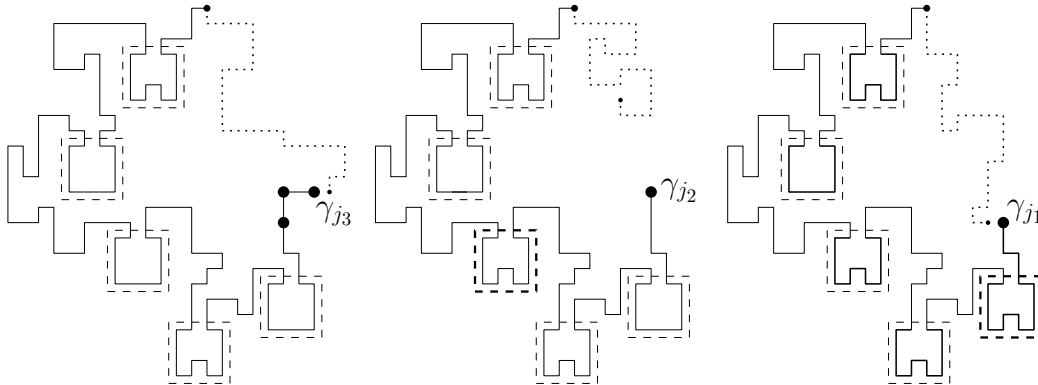


FIGURE 5. In this figure, we explain in outline the method. Given $n \in 2\mathbb{N} + 1$, the index $\ell \in [n/4, 3n/4]$ has been fixed so that the bound (4.2) holds. As we have seen, the vast majority of indices in this range satisfy this bound. This means that when we draw a length $n + 1$ polygon γ and mark with a black dot each vertex γ_j , $n/4 \leq j \leq 3n/4$, with the property that $\mathbb{W}_n^0(\Gamma \text{ closes } |\Gamma^1| = j, \Gamma^1 = \gamma_{[0,j]}) \geq n^{-1/2+2\varepsilon}$, most such γ will appear with black dots in most of the available spots. The left-hand sketch represents such a typical γ and three of its many black dots. The pointed second part shows a sample of the law $\mathbb{W}_n^0(\cdot \mid |\Gamma^1| = j_k, \Gamma^1 = \gamma_{[0,j_k]})$, with $k = 3$, one that happens to close $\gamma_{[0,j_3]}$. The second part being sampled has length $n - j_3$. If we sample instead this law with $k = 2$, then the second part has a greater length, $n - j_2$, which equals $n - j_3 + 2$ in this instance. However, to construct a charming snake, we want this length to stay the same as we move from one black dot to the next. Pattern exchange is the mechanism that achieves this. By turning one type I pattern into a pattern of type II , we push two units of length into the first part, so that, in the middle sketch, the random second part has the original length $n - j_3$. The process is iterated in the right-hand sketch. The first part is akin to a belayer who takes in rope, storing it in accumulating type II patterns, so that the second part climber maintains a constant length of rope. This process of pattern exchange can be maintained for an order of $n^{1/2}$ steps, because the Gaussian fluctuation between the two types of pattern means that the process of artificially altering pattern type does not push the system out of its rough equilibrium when the number of changes is of this order. In this way, black dots also mark charming snake terms for a snake of length of order $n^{1/2}$.

For $\gamma, \gamma' \in \text{SAP}_{n+1}$, we say that $\gamma \sim \gamma'$ if γ' may be obtained from γ by relocating the type II patterns of γ contained in the set of slots $S_1 \cup S_2$ for γ to another set of locations among these slots. The relation \sim is an equivalence relation, because the polygon γ^{empty} formed by filling all the slots of γ with type I patterns is shared by

related polygons, so that the value of $S_1 \cup S_2$ is equal for such polygons. Elements of a given equivalence class have common values of length, N_I and N_{II} , but not of N_I^i or N_{II}^i for $i \in \{1, 2\}$. We call the equivalence classes $(n+1)$ -local shells: the parameter $n+1$ appears to denote the common length of the member polygons, and the term *local* is included to indicate that members of a given class may differ only in locations that are close to the origin (in the chemical distance, along the polygon). Complementing the notation $\varsigma(\gamma)$ for the shell of γ , write $\varsigma^{\text{loc}}(\gamma)$ for the $(n+1)$ -local shell of $\gamma \in \text{SAP}_{n+1}$.

For $\varphi > 0$, write $\mathcal{G}_{n+1, \varphi}$ for the set of $(n+1)$ -local shells $\sigma \subseteq \text{SAP}_{n+1}$ such that each of the quantities $|S_1(\sigma)|$, $|S_2(\sigma)|$, N_I and N_{II} is at least $\varphi(n+1)$. Such “good” shells are highly typical if $\varphi > 0$ is small, as we now see.

Lemma 4.5. *There exist constants $c > 0$ and $\varphi > 0$ such that*

$$\mathbb{P}_{n+1}\left(\varsigma^{\text{loc}}(\Gamma) \in \mathcal{G}_{n+1, \varphi}\right) \geq 1 - e^{-cn}.$$

Proof. By Kesten’s pattern theorem [15, Theorem 1] (and Lemma 2.3, in order that we may state the result for the polygon measure), there exist constants $c > 0$ and $\varphi > 0$ such that, for any odd $n \geq d3^d$,

$$\mathbb{P}_{n+1}\left(T_I(\Gamma) \leq \varphi m\right) \leq e^{-cn} \quad \text{and} \quad \mathbb{P}_{n+1}\left(T_{II}(\Gamma) \leq \varphi m\right) \leq e^{-cn}. \quad (4.3)$$

Note that every slot in S_1 belongs to $\Gamma_{[0, (n+1)/4]}$, because such slots belong to $\Gamma_{[0, (n+1)/10 + 2T_{II}(\Gamma)]}$ and $T_{II}(\Gamma) \leq (n+1)/13$. Thus,

$$\begin{aligned} & \mathbb{P}_{n+1}\left(S_1(\Gamma) \text{ contains fewer than } \varphi(n+1) \text{ type I patterns}\right) \\ & \leq \mathbb{P}_{n+1}\left(\Gamma_{[0, (n+1)/4]} \text{ contains fewer than } \varphi(n+1) \text{ type I patterns}\right). \end{aligned}$$

There exist constants $\varphi, c > 0$ such that

$$\begin{aligned} & \mathbb{P}_{n+1}\left(\Gamma_{[0, (n+1)/4]} \text{ contains fewer than } \varphi(n+1) \text{ type I patterns}\right) \\ & \leq \frac{c(n+1)/4 c_3(n+1)/4}{p_{n+1}} \cdot \mathbb{W}_{(n+1)/4}\left(\Gamma \text{ contains fewer than } \varphi(n+1) \text{ type I patterns}\right) \\ & \leq e^{(2c_{\text{HW}}+c)\sqrt{n}} \cdot \mathbb{W}_{(n+1)/4}\left(\Gamma \text{ contains fewer than } \varphi(n+1) \text{ type I patterns}\right) < e^{-cn}, \end{aligned}$$

where the second inequality comes from the Hammersley-Welsh bound (2.1) and Lemma 2.3, and the third from (4.3) (with a relabelling of $c > 0$). Thus, $\mathbb{P}_{n+1}(N_I^1 \leq \varphi(n+1)) \leq e^{-cn}$. The same holds for the quantity N_{II}^1 . Considering $\Gamma_{[3(n+1)/4, n+1]}$ in place of $\Gamma_{[0, (n+1)/4]}$, the same conclusion may be reached about N_I^2 and N_{II}^2 . It follows that

$$\mathbb{P}_{n+1}\left(\min\{|S_1|, |S_2|, N_I, N_{II}\} < \varphi(n+1)\right) < 4e^{-cn}.$$

This completes the proof. \square

In the next lemma, we see how, for any $\gamma \in \mathcal{G}_{n+1, \varphi}$, the mixing of patterns that occurs when an element of $\zeta^{\text{loc}}(\gamma)$ is realized involves an asymptotically Gaussian fluctuation in the pattern number N_I^1 . The statement and proof are minor variations of those of [5, Lemma 3.5].

Lemma 4.6. *For any $\varphi > 0$, there exists $c > 0$ and $N \in \mathbb{N}$ such that, for $n \geq N$ and $\gamma \in \mathcal{G}_{n+1, \varphi}$,*

(1) *if $k \in \mathbb{N}$ satisfies $\left| k - \frac{T_I |S_1|}{|S_1| + |S_2|} \right| \leq n^{1/2} (\log n)^{1/4}$, then*

$$\mathbb{P}_{n+1} \left(N_I^1(\Gamma) = k \mid \Gamma \in \zeta^{\text{loc}}(\gamma) \right) \geq n^{-1/2} \exp \{ -c (\log n)^{1/2} \};$$

(2) *and, for any $g \in [1/8, 1/2]$,*

$$\mathbb{P}_{n+1} \left(\left| N_I^1(\Gamma) - \frac{N_I |S_1|}{|S_1| + |S_2|} \right| \geq n^{1/2} (\log n)^g \mid \Gamma \in \zeta^{\text{loc}}(\gamma) \right) \leq \exp \{ -c (\log n)^{2g} \}.$$

Proof. Fix a choice of γ in $\mathcal{G}_{n+1, \varphi}$. If Γ is distributed according to \mathbb{P}_{n+1} conditionally on $\Gamma \in \zeta^{\text{loc}}(\gamma)$, then N_I type I patterns and N_{II} type II patterns are distributed uniformly in the slots of $S_1 \cup S_2$. Thus, for $k \in \{0, \dots, |S_1|\}$,

$$\mathbb{P}_{n+1} \left(N_I^1(\Gamma) = k \mid \Gamma \in \zeta^{\text{loc}}(\gamma) \right) = \frac{\binom{|S_1|}{k} \binom{|S_2|}{N_I - k}}{\binom{|S_1| + |S_2|}{N_I}}. \quad (4.4)$$

Write $m = |S_1| + |S_2|$, $|S_1| = \alpha m$ and $N_I = \beta m$. By assumption $\alpha, \beta \in [\varphi, 1 - \varphi]$ and $m \geq 2\varphi(n + 1)$. Let $Z = \frac{N_I^1}{\alpha \beta m} - 1$. Under $\mathbb{P}_{n+1}(\cdot \mid \Gamma \in \zeta^{\text{loc}}(\gamma))$, Z is a random variable of mean 0, such that $\alpha \beta (1 + Z)m \in \mathbb{Z} \cap [0, \min\{|S_1|, T_I\}]$.

First, we investigate the case where Z is close to its mean. By means of a computation which uses Stirling's formula and (4.4), we find that

$$\mathbb{P}_{n+1} \left(Z = z \mid \Gamma \in \zeta^{\text{loc}}(\gamma) \right) = (1 + o(1)) \frac{\exp \left(-\frac{\alpha \beta}{2(1-\alpha)(1-\beta)} m z^2 \right)}{\sqrt{2\pi \alpha \beta (1-\alpha)(1-\beta) m}}, \quad (4.5)$$

where $o(1)$ designates a quantity tending to 0 as n tends to infinity, uniformly in the acceptable choices of γ , S_1 , S_2 and z , with $|z| \leq \frac{2n^{1/2} (\log n)^{1/2+\varepsilon}}{\alpha \beta m}$. We have obtained Lemma 4.6(1).

We now turn to the deviations of Z from its mean. From (4.4), one can easily derive that $\mathbb{P}_{n+1}(Z = z \mid \Gamma \in \zeta^{\text{loc}}(\gamma))$ is unimodal in z with maximum at the value closest to 0 that Z may take. (We remind the reader that Z takes values in $\frac{1}{\alpha \beta m} \mathbb{Z} - 1$, which contains 0 only if $\alpha \beta m \in \mathbb{Z}$.) The asymptotic equality (4.5) thus implies the existence of constants $c_0, c_1 > 0$ depending only on φ such that, for $|z| \geq \frac{n^{1/2} (\log n)^g}{\alpha \beta m}$ and n large enough,

$$\mathbb{P}_{n+1} \left(Z = z \mid \Gamma \in \zeta^{\text{loc}}(\gamma) \right) \leq c_1^{-1} n^{-1/2} \exp \{ -c_0 (\log n)^{2g} \}; \quad (4.6)$$

while for given $\varepsilon > 0$, $|z| \geq \frac{n^{1/2}(\log n)^{1/2+\varepsilon}}{\alpha\beta m}$ and n large enough,

$$\mathbb{P}_{n+1}\left(Z = z \mid \Gamma \in \zeta^{\text{loc}}(\gamma)\right) \leq c_1^{-1} n^{-1/2} \exp\{-c_0(\log n)^{1+2\varepsilon}\} \leq n^{-2}. \quad (4.7)$$

Since T_I^1 takes no more than $n+1$ values, (4.6) and (4.7) imply Lemma 4.6(2). \square

4.3. Mixing patterns by a random resampling. Consider a random resampling experiment whose law we will denote by P_{res} . First an input polygon Γ^{in} is sampled according to the law \mathbb{P}_{n+1} . Then the contents of the slots in $S_1 \cup S_2$ are forgotten and independently resampled to form an output polygon $\Gamma^{\text{out}} \in \text{SAP}_{n+1}$. That is, given Γ^{in} , Γ^{out} is chosen uniformly among the set of polygons $\gamma \in \text{SAP}_{n+1}$ for which $\gamma \in \zeta^{\text{loc}}(\Gamma^{\text{in}})$. Explicitly, if there are j type II patterns among k slots in $S_1 \cup S_2$ in Γ^{in} (so that $k \geq j$), the polygon Γ^{out} is formed by choosing uniformly at random a subset of cardinality j of these k slots and inserting type II patterns into the chosen slots.

Note the crucial property that Γ^{out} under P_{res} has the law \mathbb{P}_{n+1} : the resampling experiment holds the length $n+1$ random polygon at equilibrium. We mention in passing that a basic consequence of this resampling is a delocalization of the walk midpoint.

Proposition 4.7. [5, Proposition 1.3] *Let $d \geq 2$. There exists $C > 0$ such that, for $m \in \mathbb{N}$,*

$$\sup_{x \in \mathbb{Z}^d} \mathbb{W}_m\left(\Gamma_{\lfloor m/2 \rfloor} = x\right) \leq Cm^{-1/2}.$$

It may be instructive to consider how to prove this result using the resampling experiment (or in fact a similar one involving walks rather than polygons) and Lemma 4.5; a proof using such an approach is given in [5, Section 3.2].

Any element $\phi \in \zeta^{\text{loc}}(\gamma)$ begins by tracing a journey over the region where slots in S_1 may appear, from the origin to $\gamma_{(n+1)/10}^{\text{empty}}$; it then follows its *middle section*, the trajectory of γ from $\gamma_{(n+1)/10}^{\text{empty}}$ until $\gamma_{l'-(n+1)/10}^{\text{empty}}$ (where $l' = l^{\text{empty}}$); and it ends by moving from this vertex back to the origin, through the territory of slots in S_2 . Note that, in traversing the middle section, ϕ is exactly following a sub-walk of γ , because no pattern changes have been made to this part of γ . The timing of the middle section of this trajectory is advanced or retarded according to how many type II patterns are placed in the slots in S_1 . Each extra such pattern retards the schedule by two units. When ϕ has the *minimum* possible number $m := \max\{0, T_{II}(\gamma) - |S_2|\}$ of type II patterns in the slots of S_1 , the middle section is traversed by ϕ as early as possible, the journey taking place during $[(n+1)/10 + 2m, n - 2(T_{II}(\gamma) - m) - (n+1)/10]$. When ϕ has the maximum possible number $M := \min\{|S_1|, T_{II}(\gamma)\}$ of type II patterns in the slots of S_1 , this traversal occurs as late as possible, during $[(n+1)/10 + 2M, (n+1) - (n+1)/10 - 2(T_{II}(\gamma) - M)]$. Since $M \leq |S_1| \leq (n+1)/13$ and $T_{II}(\gamma) - m \leq |S_2| \leq (n+1)/13$, ϕ_j necessarily lies in the middle section whenever $j \in [(n+1)/10 + 2(n+1)/13, n +$

$1 - (n + 1)/10 - 2(n + 1)/13]$. Since the snake method parameter ℓ has been set to belong to the interval $[n/4, 3n/4]$, we see that ϕ_j always lies in the middle section whenever $j \in [\ell - n^{1/2}, \ell + n^{1/2}]$.

Taking $\gamma \in \text{SAP}_{n+1}$ and conditioning P_{res} on $\Gamma^{\text{in}} = \gamma$, note that the mean number of type I patterns that end up in the slots in S_1 under Γ^{out} is given by $T_I(\gamma) \cdot \frac{|S_1(\gamma)|}{|S_1(\gamma)| + |S_2(\gamma)|}$, because this expression is the product of the number of type I patterns that are redistributed and the proportion of the available slots that lie in S_1 .

Consider now a polygon $\phi \in \zeta^{\text{loc}}(\gamma)$ that achieves as closely as possible the mean value for the number of type I patterns among the slots in S_1 : that is, $T_I^1(\phi)$ equals $\lfloor T_I(\gamma) \cdot \frac{|S_1(\gamma)|}{|S_1(\gamma)| + |S_2(\gamma)|} \rfloor$. As we have noted, ϕ_ℓ is always reached during the middle section of ϕ 's three-stage journey. Define the *middle index* $l_{\text{mid}} = l_{\text{mid}}(\gamma)$ so that $\phi_\ell = \gamma_{l_{\text{mid}}}$. Note that, given γ and this value of $T_I^1(\phi)$, the value of this index is independent of the choice of ϕ .

4.4. Snakes of walks with high closing probability are typical. Recall that the index parameter ℓ (and also n) were fixed in Subsection 4.1. Moving towards the proof of Proposition 4.1, we take $\gamma \in \text{SAP}_{n+1}$ and define $\text{NoCharm}(\gamma)$ to be the set

$$\left\{ j \in (\ell - 2\mathbb{N}) \cap [l_{\text{mid}}(\gamma) - 2n^{1/2}(\log n)^{1/4}, l_{\text{mid}}(\gamma) + 2n^{1/2}(\log n)^{1/4}] : \right. \\ \left. \gamma \text{ is not } (1/2 - 2\varepsilon, n, \ell)\text{-charming at } j \right\}.$$

Henceforth in this proof, charming will mean $(1/2 - 2\varepsilon, n, \ell)$ -charming. (The parity constraint that $\ell - j$ is even is applied above because the walks of length $j + n - \ell$ considered in the definition of charming at index j must be of odd length if they are to close.)

We have seen that Γ_ℓ^{out} is visited by Γ^{out} during its middle section, when it is traversing a subpath of Γ^{in} unchanged by pattern mixing. For this reason, we may define a random variable L under P_{res} by setting $\Gamma_\ell^{\text{out}} = \Gamma_L^{\text{in}}$.

We now state a key property of the resampling procedure.

Lemma 4.8. *The events $\{\Gamma^{\text{out}} \text{ is charming at } \ell\}$ and $\{\Gamma^{\text{in}} \text{ is charming at } L\}$ coincide.*

Proof. Note that the shells of $\zeta(\Gamma_{[0,L]}^{\text{in}})$ and $\zeta(\Gamma_{[0,\ell]}^{\text{out}})$ coincide, because $\Gamma_{[0,\ell]}^{\text{out}}$ may be obtained from $\Gamma_{[0,L]}^{\text{in}}$ by modifying the I/II -status of some of its slots (these being certain slots in S_1). Thus, Lemma 4.4 implies the statement. \square

Lemma 4.9.

$$P_{\text{res}}\left(|\text{NoCharm}(\Gamma^{\text{in}})| \geq n^{1/2-\varepsilon/6}\right) \leq n^{-\varepsilon/6}.$$

Proof. Choosing $\varphi > 0$ small enough and abbreviating $\mathcal{G} = \mathcal{G}_{n+1, \varphi}$, Lemma 4.6(1) implies that, for each $\gamma \in \mathcal{G}$ and $k \in [-n^{1/2}(\log n)^{1/4}, n^{1/2}(\log n)^{1/4}]$,

$$P_{\text{res}}\left(L = \text{l}_{\text{mid}}(\Gamma^{\text{in}}) + 2k \mid \Gamma^{\text{in}} = \gamma\right) \geq n^{-1/2} \exp\{-c(\log n)^{1/2}\}.$$

Thus, again taking any $\gamma \in \mathcal{G}$,

$$\begin{aligned} & P_{\text{res}}\left(\Gamma^{\text{out}} \text{ is not charming at } \ell \mid \Gamma^{\text{in}} = \gamma\right) \\ & \geq \sum_{k=-n^{1/2}(\log n)^{1/4}}^{n^{1/2}(\log n)^{1/4}} P_{\text{res}}\left(\Gamma^{\text{out}} \text{ is not charming at } \ell, L = \text{l}_{\text{mid}}(\Gamma^{\text{in}}) + 2k \mid \Gamma^{\text{in}} = \gamma\right) \\ & \geq \sum_{k=-n^{1/2}(\log n)^{1/4}}^{n^{1/2}(\log n)^{1/4}} P_{\text{res}}\left(\gamma \text{ is not charming at } \text{l}_{\text{mid}}(\Gamma^{\text{in}}) + 2k, L = \text{l}_{\text{mid}}(\Gamma^{\text{in}}) + 2k \mid \Gamma^{\text{in}} = \gamma\right) \\ & \geq n^{-1/2} \exp\{-c(\log n)^{1/2}\} \sum_{k=-n^{1/2}(\log n)^{1/4}}^{n^{1/2}(\log n)^{1/4}} \mathbb{1}_{\gamma \text{ is not charming at } \text{l}_{\text{mid}}(\gamma) + 2k} \\ & \geq n^{-1/2} \exp\{-c(\log n)^{1/2}\} \cdot |\text{NoCharm}(\gamma)|, \end{aligned}$$

where the second inequality made use of Lemma 4.8.

Averaging over such γ , we find that

$$\begin{aligned} & P_{\text{res}}\left(\Gamma^{\text{out}} \text{ is not charming at } \ell \mid \Gamma^{\text{in}} \in \mathcal{G}\right) \\ & \geq cn^{-1/2} \exp\{-c(\log n)^{1/2}\} \cdot E_{\text{res}}\left[|\text{NoCharm}(\Gamma^{\text{in}})| \mid \Gamma^{\text{in}} \in \mathcal{G}\right], \end{aligned}$$

where E_{res} denotes the expectation associated with the law P_{res} .

Note that

$$\begin{aligned} & P_{\text{res}}\left(\Gamma^{\text{out}} \text{ is not charming at } \ell \mid \Gamma^{\text{in}} \in \mathcal{G}\right) \\ & \leq 2P_{n+1}\left(\Gamma \text{ is not charming at } \ell\right) = 2P_{n+1}\left(\Gamma_{[0, \ell]} \notin \text{HighFirst}_{\ell, n}^{1/2-2\varepsilon}\right) \leq 2n^{-\varepsilon}, \end{aligned}$$

where in the first inequality we use that Γ^{out} under P_{res} has the law P_{n+1} , and then apply Lemma 4.5, to find that $P_{\text{res}}(\Gamma^{\text{in}} \in \mathcal{G}) = P_{n+1}(\Gamma \in \mathcal{G}) \geq 1 - e^{-cn} \geq 1/2$. The final inequality above used Lemma 4.2.

Thus,

$$E_{\text{res}}\left[|\text{NoCharm}(\Gamma^{\text{in}})| \mid \Gamma^{\text{in}} \in \mathcal{G}\right] \leq n^{1/2-\varepsilon/2}.$$

We find that $E_{\text{res}}|\text{NoCharm}|$ is at most

$$\begin{aligned} & E_{\text{res}}\left[|\text{NoCharm}(\Gamma^{\text{in}})| \mid \Gamma^{\text{in}} \in \mathcal{G}\right] + \left(2n^{1/2}(\log n)^{1/4} + 1\right) P_{\text{res}}\left(\Gamma^{\text{in}} \notin \mathcal{G}\right) \\ & \leq n^{1/2-\varepsilon/2} + \left(2n^{1/2}(\log n)^{1/4} + 1\right) e^{-cn} \leq n^{1/2-\varepsilon/3}. \end{aligned}$$

Applying Markov's inequality yields Lemma 4.9. \square

Proof of Proposition 4.1. Lemma 4.6(2) implies that

$$P_{n+1}\left(\ell \in [l_{\text{mid}}(\Gamma) - n^{1/2}(\log n)^{1/4}, l_{\text{mid}}(\Gamma) + n^{1/2}(\log n)^{1/4}] \mid \Gamma \in \mathcal{G}_{n,\varphi}\right)$$

is at least $1 - \exp\{-c(\log n)^{1/2}\}$. Note that the interval centred on $l_{\text{mid}}(\Gamma)$ considered here is shorter than its counterpart in the definition of $\text{NoCharm}(\gamma)$ for $\gamma \in \text{SAP}_{n+1}$. Applying Lemmas 4.5 and 4.9, we find that

$$P_{n+1}\left(\#\left\{j \in [\ell - n^{1/2}(\log n)^{1/4}, \ell + n^{1/2}(\log n)^{1/4}] : \Gamma \text{ is not charming at } j\right\} \geq n^{1/2-\varepsilon/6}\right) \leq n^{-\varepsilon/6} + e^{-c(\log n)^{1/2}} + e^{-cn}.$$

When the complementary event occurs, Γ is charming for at least one-quarter of the indices in $[\ell - n^{1/2}, \ell]$, so that $\Gamma_{[0,\ell]}$ is an element of $\text{CS}_{1/2,0}^{1/2-2\varepsilon,\ell,n}$. (We write one-quarter rather than one-half here, because one-half of such indices are inadmissible due to their having the wrong parity.) \square

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DEPARTMENTS OF MATHEMATICS AND STATISTICS, U.C. BERKELEY, BERKELEY, CA, 94720-3840, U.S.A.

E-mail address: `alanmh@stat.berkeley.edu`