## Chapter 7: Likelihood Inference

Carlos Enrique Améndola Cerón

Introduction to Algebraic Statistics Course

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- (Def 5.3.5) Likelihood function for a model  $M_{\Theta}$  with data D:  $L(\theta|D)$ (=  $p_{\theta}(D)$  or  $f_{\theta}(D)$ ).
- MLE  $\hat{\theta}$  maximizes the (log-)likelihood function:

$$\hat{ heta} = rg\max_{ heta \in \Theta} \, \ell( heta | D)$$

- (Def 7.1.1) The score equations are obtained by setting the gradient of the log-likelihood to zero:  $\frac{\partial}{\partial \theta_i} \ell(\theta|D) = 0$  for i = 1, ..., d.
- In the discrete case  $p: \Theta \to \Delta_{r-1}$ : for i.i.d. data  $X^{(1)}, \ldots, X^{(n)}$  summarized by the vector of counts  $u \in \mathbb{N}^r$ , we have

$$\ell(\theta|u) = \sum_{j=1}^{r} u_j \log p_j(\theta).$$

•  $\ell(\theta|u) = \sum_{j=1}^{r} u_j \log p_j(\theta)$ , hence score equations are *rational*:

$$\sum_{j=1}^{r} \frac{u_j}{p_j} \frac{\partial p_j}{\partial \theta_i}(\theta) = 0 \quad i = 1 \dots, d.$$

### Theorem (Thm 7.1.2, Def 7.1.4)

Let  $p: \Theta \to \Delta_{r-1}$ . For generic data, the number of (complex) solutions to the score equations is independent of u. We call this the ML degree of the parametric discrete statistical model  $M_{\Theta} \subset \Delta_{r-1}$ .

- ML degree measures the complexity of the ML estimation problem.
- ML degree is  $1 \iff$  the MLE is a rational function of the data.

### Example (Twisted Cubic Model)

$$p( heta) = (s, s heta, s heta^2, s heta^3) \subset \Delta_3 \subset \mathbb{R}^4.$$

where  $s = \frac{1}{1+\theta+\theta^2+\theta^3}$ . Sample size  $n = u_0 + u_1 + u_2 + u_3$ . We have

$$L(\theta|u) = s^{u_0}(s\theta)^{u_1}(s\theta^2)^{u_2}(s\theta^3)^{u_3}$$
  
=  $s^{u_0+u_1+u_2+u_3}\theta^{u_1+2u_2+3u_3}$ 

$$\ell(\theta|u) = n\log s + (u_1 + 2u_2 + 3u_3)\log\theta$$

The score equation is:

$$0 = \frac{\partial \ell}{\partial \theta} = -ns(1+2\theta+3\theta^2) + (u_1+2u_2+3u_3)\frac{1}{\theta}$$

Thus  $3n\theta^3 + 2n\theta^2 + n\theta - (u_1 + 2u_2 + 3u_3)s^{-1} = 0$  and we arrive at

$$3(n - u_3)\theta^3 + 2(n - u_2)\theta^2 + (n - u_1)\theta - (u_1 + 2u_2 + 3u_3) = 0$$

The ML degree is 3.

• Recall (Prop 5.3.7) the Gaussian model log-likelihood  $\ell(\mu, \Sigma | \bar{X}, S)$ :

$$-\frac{n}{2}(\log \det \Sigma + m \log 2\pi) - \frac{n}{2} \operatorname{tr}(S\Sigma^{-1}) - \frac{n}{2}(\bar{X} - \mu)^T \Sigma^{-1}(\bar{X} - \mu).$$

### Example (Prop 7.1.6)

Let  $\Theta = \Theta_1 \times Id_m \subset \mathbb{R}^m \times PD_m$  for a Gaussian statistical model. Then the maximum likelihood estimation for  $\Theta$  is equivalent to the least-squares point on  $\Theta_1$ . In this case, ML degree = # critical points of  $||\bar{X} - \mu||_2^2$ , known as the ED degree of  $\Theta_1$ .

• (Prop 7.1.9) Let  $\Theta = \mathbb{R}^m \times \Theta_2 \subset \mathbb{R}^m \times PD_m$  for a Gaussian statistical model. Then ML estimation gives  $\hat{\mu} = \bar{X}$  and reduces to maximizing  $-\frac{n}{2} \log \det \Sigma - \frac{n}{2} \operatorname{tr}(S\Sigma^{-1})$ .

### Example (Ex 7.1.11 Gaussian Marginal Independence)

Let  $\Theta = \mathbb{R}^m \times \Theta_2$  where  $\Theta_2 = \{\Sigma \in PD_4 | \sigma_{12} = \sigma_{21} = 0, \sigma_{34} = \sigma_{43} = 0\}$ . The marginal independence constraints are  $X_1 \perp \!\!\!\perp X_2$  and  $X_3 \perp \!\!\!\perp X_4$ . The ML degree is found to be 17.

## Definition (ML degree of a variety)

Let  $V \subset \mathbb{P}^{r-1}$  be an irreducible projective variety over  $\mathbb{C}$ ,  $u \in \mathbb{N}^r$  and

$$L_{u}(p) = \frac{p_{1}^{u_{1}}p_{2}^{u_{2}}\cdots p_{r}^{u_{r}}}{(p_{1}+\cdots+p_{r})^{u_{1}+\cdots+u_{r}}}$$

ML degree of V is the number of (complex) critical points for generic u of  $L_u(p)$  on  $V_{reg} \setminus \mathcal{H}$ , where  $\mathcal{H} = \{p \in \mathbb{P}^{r-1} : p_1 \cdots p_r(p_1 + \cdots + p_r) = 0\}$ .

- If  $I(V) = \langle f_1, f_2, \dots, f_k \rangle$ , use Lagrange multipliers to optimize L.
- (Thm 7.2.9) Huh (2013): the ML degree of a smooth very affine variety (of the form V ∩ (ℂ\*)<sup>r</sup> where V ⊂ ℂ<sup>r</sup> variety) is ±χ<sub>top</sub>(·).
- (Theorem 7.2.13) Huh (2014): Characterization of ML degree 1 varieties as *A*-discriminants [GKZ] (via *Horn uniformization*).

# ML in Exponential Families

### Theorem (Prop 7.3.7)

Exponential family  $p_{\theta}(x) = h(x) \exp(\langle \theta, T(x) \rangle - A(\theta))$  with sufficient statistics T(x), log-partition function  $A(\theta) = \log \int_{\mathcal{X}} h(x) \exp(\langle \theta, T(x) \rangle)$ Then

$$\frac{\partial}{\partial \theta_i} A(\theta) = \mathbb{E}_{\theta}[T_i(X)] \quad \text{and} \quad \frac{\partial^2}{\partial \theta_i \theta_j} A(\theta) = \operatorname{Cov}_{\theta}[T_i(X), T_j(X)]$$

### Corollary (Cor 7.3.8)

The likelihood function for an exponential family is strictly concave. The MLE (if it exists) is the unique solution to the equation

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\mathbb{E}_{\theta}[T(X)] = T(x)
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where x denotes the data vector.

## Corollary (Birch's Theorem, Cor 7.3.9)

The MLE in the log-linear model  $\mathcal{M}_{A,h}$  given the data u is the unique solution, if it exists, to the equations

 $Au = n A\hat{p}$  and  $\hat{p} \in \mathcal{M}_{A,h}$ 

Inspires algorithms for computing MLE: Iterative Proportional Scaling (IPS)

### Corollary (Cor 7.3.10)

Let  $X^{(1)}, \ldots, X^{(n)} \in \mathbb{R}^m$  i.i.d. samples from the Gaussian exponential family parametrized by  $(\mu, \Sigma) \in \mathbb{R}^m \times \mathcal{M}_{L^{-1}}$  (L linear space such that  $L \cap PD_m \neq \emptyset$ ). The MLE is  $(\bar{X}, \hat{S})$  where  $\hat{S}$  is the unique solution (if it exists) to the equations

$$\pi(S) = \pi(\hat{S})$$
 and  $\hat{S} \in \mathcal{M}_{L^{-1}}$ 

where  $\pi$  denotes the orthogonal projection onto L.

Let  $\mathcal{M}$  be the model of binomial random variables  $Bin(2, \theta)$ :

$$\mathcal{M} = \{((1- heta)^2, 2 heta(1- heta), heta^2) \in \Delta_2 \,|\, heta \in (0,1)\}$$

- What is the ML degree of  $\mathcal{M}$ ?
- Compute the MLE  $\hat{\theta}$  for the two data points u = (8, 6, 5) and v = (4, 20, 8). Interpret your results.